

Lecture note 27. More theory of martingales. Discrete time.

**Microtheorem 27.1.** Let  $Y_t, t \in T$ , be a martingale with respect to a stochastic process  $X_t, t \in T$ . Let  $\tau$  be a stopping time taking finitely many values  $t_1 < t_2 < \dots < t_n$  (all in  $T$ , not  $+\infty$ ).

Then for an arbitrary  $t_* \in T$

$$E(Y_\tau) = E(Y_{t_*}). \quad (27.1)$$

By the way, I did not mention this before, but the expectation (unconditional) of a martingale  $Y_t$  does not depend on  $t$ : for  $t < s$  we have:

$$E(Y_s) = E(E(Y_s \| X_u, u \leq t)) = E(Y_t). \quad (27.2)$$

**Proof of Microtheorem 27.1.** We have:

$$E(Y_\tau) = \sum_{i=1}^n E(I_{\{\tau=t_i\}}(\omega) \cdot Y_\tau) = \sum_{i=1}^n E(I_{\{\tau=t_i\}}(\omega) \cdot Y_{t_i}). \quad (27.3)$$

For  $i > 1$  the event  $\{\tau = t_i\}$  whose indicator is used in this formula is equal to the difference of two:

$$\{\tau = t_i\} = \{\tau \leq t_i\} \setminus \{\tau \leq t_{i-1}\}; \quad (27.4)$$

and for  $i = 1$  it is just

$$\{\tau = t_1\} = \{\tau \leq t_1\}. \quad (27.5)$$

So the  $i$ -th summand in (27.3) can be written as

$$E((I_{\{\tau \leq t_i\}}(\omega) - I_{\{\tau \leq t_{i-1}\}}(\omega)) \cdot Y_{t_i}) \quad (27.6)$$

for  $i > 1$ , and as

$$E(I_{\{\tau \leq t_1\}}(\omega) \cdot Y_{t_1}) \quad (27.7)$$

for  $i = 1$ . So we get:

$$\begin{aligned} E(Y_\tau) &= E(I_{\{\tau \leq t_1\}}(\omega) \cdot Y_{t_1}) - E(I_{\{\tau \leq t_1\}}(\omega) \cdot Y_{t_2}) + E(I_{\{\tau \leq t_2\}}(\omega) \cdot Y_{t_2}) \\ &\quad - E(I_{\{\tau \leq t_2\}}(\omega) \cdot Y_{t_3}) + \dots + E(I_{\{\tau \leq t_{n-1}\}}(\omega) \cdot Y_{t_{n-1}}) \\ &\quad - E(I_{\{\tau \leq t_{n-1}\}}(\omega) \cdot Y_{t_n}) + E(I_{\{\tau \leq t_n\}}(\omega) \cdot Y_{t_n}) \\ &= \sum_{i=1}^{n-1} E(I_{\{\tau \leq t_i\}}(\omega) \cdot (Y_{t_i} - Y_{t_{i+1}})) + E(I_{\{\tau \leq t_n\}}(\omega) \cdot Y_{t_n}). \end{aligned} \quad (27.8)$$

We represent the  $i$ -th summand ( $i < n$ ) as the expectation of a conditional expectation:

$$E(I_{\{\tau \leq t_i\}}(\omega) \cdot (Y_{t_{i+1}} - Y_{t_i})) = E(E(I_{\{\tau \leq t_i\}}(\omega) \cdot (Y_{t_i} - Y_{t_{i+1}})) \| X_u, u \leq t_i)). \quad (27.9)$$

The event  $\{\tau \leq t_i\}$  is represented as  $\{(X_u, u \leq t_i) \in C\}$  (i. e., is determined by the values of  $X_u$  with  $u \leq t_i$ ), so its indicator can be taken from under the sign of the conditional expectation:

$$E(I_{\{\tau \leq t_i\}}(\omega) \cdot (Y_{t_i} - Y_{t_{i+1}})) = E(I_{\{\tau \leq t_i\}}(\omega) \cdot E(Y_{t_i} - Y_{t_{i+1}} \| X_u, u \leq t_i)); \quad (27.10)$$

and by the martingale property, the conditional expectation is equal to 0.

The last summand in (27.8) is equal to  $E(Y_{t_n})$ , because the event  $\{\tau \leq t_n\}$  is the whole  $\Omega$ , and its indicator is identically 1. By the way,  $E(Y_{t_n})$  is equal to  $E(Y_{t_*})$  for any other  $t_* \in T$ .

The same way we prove

**Microtheorem 27.2.** *Let  $Y_t, t \in T$ , be a submartingale with respect to a stochastic process  $X_t, t \in T$ . Let  $\tau$  be a stopping time taking finitely many values  $t_1 < t_2 < \dots < t_n = t_{\max}$ .*

*Then*

$$E(Y_\tau) \leq E(Y_{t_{\max}}) \quad (27.11)$$

(for a submartingale,  $E(Y_t)$  is no longer a constant, but rather a nondecreasing function).

**Proof.** Formulas (27.8), (27.9), (27.10) still hold; but instead of saying “by the martingale property the conditional expectation is equal to 0” we say “by the submartingale property the conditional expectation is  $\leq 0$ ”.

For random times  $\tau$  that are not stopping times, the statements of Microtheorems 27.1, 27.2 are not necessarily true. E. g., if  $X_0 \equiv 0$ , and  $X_1$  is a random variables with density  $p(x)$  with  $E(X_1) = 0$ , the sequence of two random variables  $X_0, X_1$  is a martingale. Let us define the random time  $\tau$  by

$$\tau = I_{[0, \infty)}(X_1). \quad (27.12)$$

This is *not* a stopping time: to know whether  $\tau \leq 0$  (which means  $= 0$  in this case), it is not enough to watch  $X_0$  (as if there were anything to watch here!), but we need to know  $X_1$ . We have:

$$E(X_\tau) = E(I_{\{\tau=0\}} \cdot X_0 + I_{\{\tau=1\}} \cdot X_1) = E(I_{[0, \infty)}(X_1) \cdot X_1) = \int_0^\infty x \cdot p(x) dx > 0, \quad (27.13)$$

$$\text{while } E(X_1) = \int_{-\infty}^\infty x \cdot p(x) dx = 0.$$

Let  $Y_t, t \in T$ , be a random function;  $\tau$  a random variable taking values in  $T \cup \{\infty\}$ . The random function  $\hat{Y}_t, t \in T$ , obtained by stopping  $Y_t$  at the time  $\tau$  is defined as

$$\hat{Y}_t = \begin{cases} Y_t, & t < \tau, \\ Y_\tau, & t \geq \tau. \end{cases} \quad (27.14)$$

**Microtheorem 27.3.** *If  $Y_t, t \in T$ , is a martingale (a submartingale), and  $\tau$  a stopping time taking finitely many values, the stopped random function  $\hat{Y}_t, t \in T$ , is also a martingale (a submartingale).*

**Proof.** We have to prove that for  $t < s$

$$E(\hat{Y}_s - \hat{Y}_t \| X_u, u \leq t) = 0 \quad (\text{or } \geq 0). \quad (27.15)$$

Without restriction of generality, we can assume that  $t$  and  $s$  are among the values  $t_i$  that the stopping time  $\tau$  takes:  $t = t_m, s = t_k$ . If not, we can add  $t$  and  $s$  to these values, and state that  $\tau = t$  or  $s$  only with probability 0.

So we have to prove that

$$E(\hat{Y}_{t_k} - \hat{Y}_{t_m} \| X_u, u \leq t) = \sum_{i=m+1}^k E(\hat{Y}_{t_i} - \hat{Y}_{t_{i-1}} \| X_u, u \leq t) = 0 \quad (\text{or } \geq 0). \quad (27.16)$$

It's enough to prove that every summand is  $= 0$  (or  $\geq 0$ ). Using the property 3) of conditional expectations, we get:

$$E(\hat{Y}_{t_i} - \hat{Y}_{t_{i-1}} \| X_u, u \leq t) = E(E(\hat{Y}_{t_i} - \hat{Y}_{t_{i-1}} \| X_u, u \leq t_{i-1}) \| X_u, u \leq t). \quad (27.17)$$

Again, it is enough to prove that

$$E(\hat{Y}_{t_i} - \hat{Y}_{t_{i-1}} \| X_u, u \leq t_{i-1}) = 0 \quad (\text{or } \geq 0). \quad (27.18)$$

Formula (27.18) means that for every set  $C$  in the space of functions  $x_u, u \leq t_{i-1}$ ,

$$E(I_C(X_u, u \leq t_{i-1}) \cdot (\hat{Y}_{t_i} - \hat{Y}_{t_{i-1}})) = 0 \quad (\text{or } \geq 0). \quad (27.19)$$

We haven't used the fact that  $\tau$  is a stopping time. And this means, in particular, that there exists a set  $D$  in the space of functions  $x_u, u \leq t_{i-1}$ , such that

$$\{\tau \leq t_{i-1}\} = \{(X_u, u \leq t_{i-1}) \in D\} \quad (27.20)$$

(I seem again to run out of letters). The opposite event

$$\{\tau > t_{i-1}\} = \{(X_u, u \leq t_{i-1}) \in D^c\}, \quad (27.21)$$

where  $D^c$  is the complement to the set  $D$  in the space of functions  $x_u, u \leq t_{i-1}$ .

The expectation (27.19) can be represented as

$$\begin{aligned} & E(I_C(X_u, u \leq t_{i-1}) \cdot I_D(X_u, u \leq t_{i-1}) \cdot (\hat{Y}_{t_i} - \hat{Y}_{t_{i-1}})) \\ & + E(I_C(X_u, u \leq t_{i-1}) \cdot I_{D^c}(X_u, u \leq t_{i-1}) \cdot (\hat{Y}_{t_i} - \hat{Y}_{t_{i-1}})). \end{aligned} \quad (27.22)$$

In the first summand the second indicator is the indicator of the event  $\{\tau \leq t_{i-1}\}$ , so in the first summand  $\hat{Y}_{t_i} = \hat{Y}_{t_{i-1}} = Y_\tau$ , and this first summand is equal to 0. In the second summand we have  $\tau > t_{i-1}, \tau \geq t_i, \hat{Y}_{t_{i-1}} = Y_{t_{i-1}}, \hat{Y}_{t_i} = Y_{t_i}$ , and the second summand is equal to

$$E(I_{C \cap D^c}(X_u, u \leq t_{i-1}) \cdot (Y_{t_i} - Y_{t_{i-1}})), \quad (27.23)$$

and by the martingale (or the submartingale) property this expectation is  $= 0$  (or  $\geq 0$ ).

**Microtheorem 27.4** (the Kolmogorov-type inequality). *If  $Y_t, t \in T$ , is a nonnegative submartingale (with respect to  $X_t, t \in T$ ), and  $t_1 < t_2 < \dots < t_n = t_{\max}$  are points in  $T$ , then for every  $a > 0$*

$$P\{\max(Y_{t_1}, Y_{t_2}, \dots, Y_{t_n}) \geq a\} \leq \frac{E(Y_{t_{\max}})}{a}. \quad (27.24)$$

This is a strengthening of the following Chebyshev-type inequality:

$$P\{Y_{t_i} \geq a\} \leq \frac{E(Y_{t_i})}{a} \leq \frac{E(Y_{t_{\max}})}{a}, \quad (27.25)$$

only the Chebyshev-type inequality is about *one* random variable, and (27.24) is about the maximum of an arbitrarily large number of random variables.

What Kolmogorov started this with was a particular case:

**Microtheorem 27.5.** *Let  $X_1, X_2, \dots, X_n$  be independent random variables with  $E(X_i) = 0$  and finite variances  $\text{Var}(X_i) = \sigma_i^2$ . Then for every  $b > 0$*

$$P\{\max(|X_1|, |X_1 + X_2|, \dots, |X_1 + X_2 + \dots + X_n|) \geq b\} \leq \frac{\sigma_1^2 + \sigma_2^2 + \dots + \sigma_n^2}{b^2}. \quad (27.26)$$

The classical Kolmogorov inequality is a strengthening of the classical Chebyshev's inequality:

$$P\{|X_1 + \dots + X_i| \geq b\} \leq \frac{\text{Var}(X_1 + \dots + X_i)}{b^2} = \frac{\sigma_1^2 + \dots + \sigma_i^2}{b^2} \leq \frac{\sigma_1^2 + \dots + \sigma_n^2}{b^2}. \quad (27.27)$$

Let me show how the classical Kolmogorov's inequality (27.26) is a particular case of the Kolmogorov-type inequality (27.24). The sequence  $Y_i = X_1 + \dots + X_i, i = 1, \dots, n$ , is a martingale with respect to the sequence of the random variables  $X_i$  (see Example 23.2, generalization). By Example 24.1, the sequence  $Z_i = (X_1 + \dots + X_i)^2, i = 1, \dots, n$ , is a submartingale; and by inequality (27.24) we have:

$$P\{\max(|X_1|, |X_1 + X_2|, \dots, |X_1 + \dots + X_n|) \geq b\} = P\{\max(Z_1, Z_2, \dots, Z_n) \geq b^2\} \leq \frac{E(Z_n^2)}{b^2}, \quad (27.28)$$

which is equal to the right-hand side of (27.26).

**Proof of Microtheorem 27.4.** Let us introduce the random time

$$\tau = \begin{cases} \min\{t_i : 1 \leq i \leq n, Y_{t_i} \geq a\} & \text{if there is at least one such } t_i, \\ \infty & \text{if there is no such } t_i. \end{cases} \quad (27.29)$$

Similarly to Example 25.3,  $\tau$  is a stopping time. By Example 26.2,  $\min(\tau, t_{\max})$  is also a stopping time – and this one, taking finitely many finite values.

The event  $\{\max(Y_{t_1}, Y_{t_2}, \dots, Y_{t_n}) \geq a\}$  clearly coincides with the event

$$\{Y_{\min(\tau, t_{\max})} \geq a\}. \quad (27.30)$$

Applying a Chebyshev-type inequality to the nonnegative random variable  $Y_\tau$ , we get:

$$P\{\max(Y_{t_1}, Y_{t_2}, \dots, Y_{t_n}) \geq a\} = P\{Y_{\min(\tau, t_{\max})} \geq a\} \leq \frac{E(Y_{\min(\tau, t_{\max})})}{a}, \quad (27.31)$$

and by Microtheorem 27.2 this is  $\leq \frac{E(Y_{t_{\max}})}{a}$ .